

5

$$y = X\beta_0 + u \quad u \sim N(0, \sigma_0^2 I_n), \quad \theta = (\beta_0', \sigma_0^2)' \in \mathbb{R}^{k+1}$$

$$\text{Log-likelihood Funktion } \ell_n(\theta) = -\frac{n}{2} \ln \sigma_0^2 - \frac{1}{2\sigma_0^2} \|y - X\beta\|^2$$

$$\frac{\partial}{\partial \theta} \ell_n(\theta) = \begin{pmatrix} -\frac{1}{\sigma_0^2} (X'X\beta - X'y) \\ -\frac{n}{2\sigma_0^2} + \frac{1}{2\sigma_0^4} \|y - X\beta\|^2 \end{pmatrix}$$

$$\frac{\partial^2}{\partial \theta \partial \theta'} \ell_n(\theta) = \begin{pmatrix} -\frac{1}{\sigma_0^2} X'X & \frac{1}{\sigma_0^4} (X'X\beta - X'y) \\ \frac{1}{\sigma_0^4} (X'X\beta - X'y)' & \frac{n}{2\sigma_0^4} - \frac{1}{\sigma_0^6} \|y - X\beta\|^2 \end{pmatrix}$$

$$\begin{aligned} \Rightarrow E_{\theta_0} \left(-\frac{\partial^2}{\partial \theta \partial \theta'} \ell_n(\theta_0) \right) &= \begin{pmatrix} \frac{1}{\sigma_0^2} X'X & E_{\theta_0}(X'y - X'X\beta_0) \\ E_{\theta_0}(X'y - X'X\beta_0)' & \frac{n}{2\sigma_0^4} - \frac{1}{\sigma_0^6} E \|y - X\beta_0\|^2 \end{pmatrix} \\ &= \begin{pmatrix} \frac{1}{\sigma_0^2} X'X & 0 \\ 0 & \frac{n}{2\sigma_0^4} \end{pmatrix} = I(\theta_0) \\ &= \frac{1}{\sigma_0^2} X'X \quad 0 \\ & \quad 0 \quad \frac{n}{2\sigma_0^4} \end{pmatrix} = I(\theta_0) \\ &= \frac{n}{2\sigma_0^4} - \frac{1}{\sigma_0^6} E \|u\|^2 \\ &= \frac{n}{2\sigma_0^4} - \frac{1}{\sigma_0^6} n \sigma_0^2 = \frac{n}{2\sigma_0^4} \end{aligned}$$

$$\text{Restriktion } R\beta = r = 0: \quad g(\theta) = R\beta - r = 0 \quad R \in \mathbb{R}^{q \times k}$$

$$\Rightarrow \frac{\partial}{\partial \theta} g(\theta) = (R \ 0) = \tilde{R} \in \mathbb{R}^{q \times (k+1)}$$

$$\begin{aligned} \text{Wald-Statistik } T_W &= g(\hat{\theta}_{ML})' \left(\frac{\partial}{\partial \theta} g(\hat{\theta}_{ML})' I(\hat{\theta}_{ML})^{-1} \frac{\partial}{\partial \theta} g(\hat{\theta}_{ML}) \right)^{-1} g(\hat{\theta}_{ML}) \\ &= (R\hat{\beta} - r)' \left((R \ 0) \begin{pmatrix} \frac{1}{\hat{\sigma}_{ML}^2} X'X & 0 \\ 0 & \frac{n}{2\hat{\sigma}_{ML}^4} \end{pmatrix}^{-1} \begin{pmatrix} R \\ 0 \end{pmatrix} \right)^{-1} (R\hat{\beta} - r) \\ &= (R\hat{\beta} - r)' \left((R \ 0) \begin{pmatrix} \hat{\sigma}_{ML}^2 (X'X)^{-1} & 0 \\ 0 & \frac{2\hat{\sigma}_{ML}^4}{n} \end{pmatrix} \begin{pmatrix} R \\ 0 \end{pmatrix} \right)^{-1} (R\hat{\beta} - r) \end{aligned}$$

$$= \frac{(R\hat{\beta} - r)' (R(X'X)^{-1}R')^{-1} (R\hat{\beta} - r)}{\hat{\sigma}_{ML}^2}$$

da $\hat{\sigma}_{ML}^2 = \frac{1}{n} \hat{u}'\hat{u} = \frac{n-k}{n} \hat{\sigma}^2$ gilt:

$$T_W = \frac{(R\hat{\beta} - r)' (R(X'X)^{-1}R')^{-1} (R\hat{\beta} - r)}{\frac{n-k}{n} \hat{\sigma}^2}$$

$$= \frac{q \cdot n}{n-k} \underbrace{\frac{(R\hat{\beta} - r)' (R(X'X)^{-1}R')^{-1} (R\hat{\beta} - r)}{\hat{\sigma}^2}}_{\text{F-Statistik}} \cdot \frac{1}{q}$$

$$\begin{array}{c} n \rightarrow \infty \\ \downarrow \\ q \end{array}$$

$$= \text{F-Statistik} \sim F_{q, n-k}$$

übrigens gilt: $F_{q, n-k} \xrightarrow{n \rightarrow \infty} \chi_q^2$